Heat-treated virus inactivation rate depends strongly on treatment procedure: illustration with SARS-CoV-2

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Supplemental Material

Contents

1 Literature review

Figure S1. Selection process for literature review. Review assessed heat-treatment procedure description quality for coronavirus inactivation studies. This figure was made following the Preferred Reporting Items for Systematic Reviews and Meta-Analyses [\[1\]](#page-8-0).

2 Bayesian estimation models

2.1 Model notation

In the model notation that follows, the symbol ∼ denotes that a random variable is distributed according to a given distribution. Normal distributions are parametrized as:

Normal(mean,standard deviation)

Positive-constrained normal distributions ("Half-Normal") are parametrized as:

Half-Normal(mode,standard deviation)

2.2 Titer inference

We inferred individual titers directly from titration well data using a Poisson single-hit model. We assigned a weakly informative Normal prior to the log_{10} titers v_i (v_i is the titer for sample i measured in $\log_{10} \text{TCID}_{50}/0.1 \text{mL}$, since wells were inoculated with 0.1mL :

$$
v_i \sim \text{Normal}(2.5, 3) \tag{1}
$$

We then modeled individual positive and negative wells for sample i according to a Poisson single-hit model [\[44\]](#page-10-2). That is, for an undiluted inoculum, the number of virions that successfully infect cells within a given well is Poisson distributed with mean:

$$
\ln(2) 10^{v_i} \tag{2}
$$

The value of the mean derives from the fact that our units are $TCID_{50}$; the probability of a positive well at $v_i = 0$, i.e. 1 TCID₅₀, is equal to $1 - e^{-\ln(2) \times 1} = 0.5$.

Let Y_{idk} be a binary variable indicating whether the k^{th} well at dilution factor d (where d expressed as \log_{10} dilution factor) for sample i was positive (so $Y_{idk} = 1$ if that well was positive and 0 if it was negative). Under a single-hit process, a well will be positive as long as at least one virion successfully infects a cell.

It follows from equation [2](#page-3-3) that the conditional probability of observing $Y_{iak} = 1$ given a true underlying log_{10} titer v_i and a dilution factor d is given by:

$$
\mathcal{L}(Y_{idk} = 1 \mid v_i) = 1 - e^{-\ln(2) \times 10^{(v_i - d)}} \tag{3}
$$

This is simply the probability that a Poisson random variable with mean $\ln(2) \times 10^{(v_i-d)}$ is greater than 0. That mean is the expected number virions inoculated into the well; it derives from the fact that there are $v_i - d \log_{10} \text{TCID}_{50}$ in the dilute sample. Similarly, the conditional probability of observing $Y_{idk} = 0$ given a true underlying log_{10} titer v_i is:

$$
\mathcal{L}(Y_{idk} = 0 \mid v_i) = e^{-\ln(2) \times 10^{(v_i - d)}} \tag{4}
$$

which is the probability that the Poisson random with variable is equal to 0.

This gives us our likelihood function, assuming independence of outcomes across wells. Our inoculated doses were of volume 0.1 mL, so we incremented inferred titers by 1 to convert to units of $\log_{10}TCID_{50}/mL$.

2.3 Virus inactivation regression

Duration of virus of detectability depends not only on environmental conditions and treatment method but also initial inoculum and sampling noise. We therefore estimated the exponential decay rates of viable virus (and thus virus half-lives) using a Bayesian regression analogous to that used in [\[18,](#page-8-17) [45\]](#page-10-3). This modeling approach allowed us to account for differences in initial inoculum levels across samples as well as other sources of experimental noise. The model yields estimates of posterior distributions of viral decay rates and half-lives in the various experimental conditions – that is, estimates of the range of plausible values for these parameters given our data, with an estimate of the overall uncertainty [\[46\]](#page-10-4).

Our data consist of four different experimental conditions corresponding to four heat-treatment procedures, all at 70° C: (1) an uncovered plate of wells in a dry oven, (2) a covered plate in the oven, (3) a set of closed vials in the oven, and (4) set of closed vials in a heat block.

For each treatment, we took three samples per time point at multiple time-points.

We model each sample j for experimental condition i as starting with some true initial log₁₀ titer: v_{ij0} . At the time t_{ij} that it is sampled, it has titer v_{ij} .

We assume that viruses in experimental condition i decay exponentially at a rate λ_i over time. It follows that:

$$
v_{ij} = v_{ij0} - \lambda_i t_{ij} \tag{5}
$$

We use the direct-from-well data likelihood function described above, except that now instead of titers we estimate λ_i under the assumptions that our observed well data Y_{idk} reflect the titers v_{ij} .

We assume that each experiment i has a mean initial log₁₀ titer \bar{v}_{i0} . We model the individual initial titers v_{ij0} as normally distributed about \bar{v}_{i0} with an estimated, experiment-specific standard deviation σ_i :

$$
v_{ij0} \sim \text{Normal}(\bar{v}_{i0}, \sigma_i) \tag{6}
$$

2.4 Regression prior distributions

We placed a Normal prior on the mean initial \log_{10} titers \bar{v}_{i0} that reflects the known inocula.

$$
\bar{v}_{i0} \sim \text{Normal}(4.5, 0.5) \tag{7}
$$

We placed a Half-Normal prior on the standard deviations σ_i that allows for potentially large variation (1) log) variation about the experiment mean, as well as for less variation:

$$
\sigma_i \sim \text{Half-Normal}(0, 0.25) \tag{8}
$$

To encode prior information about virus inactivation rate in an interpretable way, we placed a Normal prior on the log half-lives $\ln(h_i)$, where $h_i = \frac{\ln(2)}{\lambda_i}$ $\frac{i(2)}{\lambda_i}$:

$$
\ln(h_i) \sim \text{Normal}(\ln(0.5), 2) \tag{9}
$$

This prior reflects that both of rapid virus inactivation and substantially slower inactivation are plausible a priori.

2.5 Predictive checks

We assessed the appropriateness of prior distribution choices using prior predictive checks and assessed goodness of fit for the estimated model using posterior predictive checks. The resultant checks are shown below.

Figure S2. Titer estimation prior check. Violin plots show distribution of simulated titers sampled from the prior predictive distribution. Points show estimated titers for each collected sample; vertical bar shows a 95% credible interval. Time-points with no positive wells for any replicate are plotted as triangles at the approximate single-replicate detection limit of the assay (LOD; denoted by a black dotted line at $10^{0.5}$ $TCID₅₀/mL$ media) to indicate that a range of sub-LOD values are plausible. Wide coverage of violins relative to datapoints show that priors are agnostic over the titer values of interest.

Figure S3. Prior predictive check for regression model. Violin plots show distribution of simulated titers sampled from the prior predictive distribution. Points show estimated titers for each collected sample; vertical bar shows a 95% credible interval. Time-points with no positive wells for any replicate are plotted as triangles at the approximate single-replicate detection LOD (denoted by a black dotted line at $10^{0.5}$ $TCID₅₀/mL$ media) to indicate that a range of sub-LOD values are plausible. Wide coverage of violins relative to datapoints show that priors are agnostic over the titer values of interest, and that the priors regard both fast and slow decay rates as possible.

Figure S4. Posterior predictive check for regression model. Violin plots show distribution of simulated titers sampled from the posterior predictive distribution. Points show estimated titers for each collected sample; vertical bar shows a 95% credible interval. Time-points with no positive wells for any replicate are plotted as triangles at the approximate single-replicate detection LOD (denoted by a black dotted line at $10^{0.5}$ TCID₅₀/mL media) to indicate that a range of sub-LOD values are plausible. Close correspondence between distribution of posterior simulated titers and estimated titers suggests the model fits the data well.

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